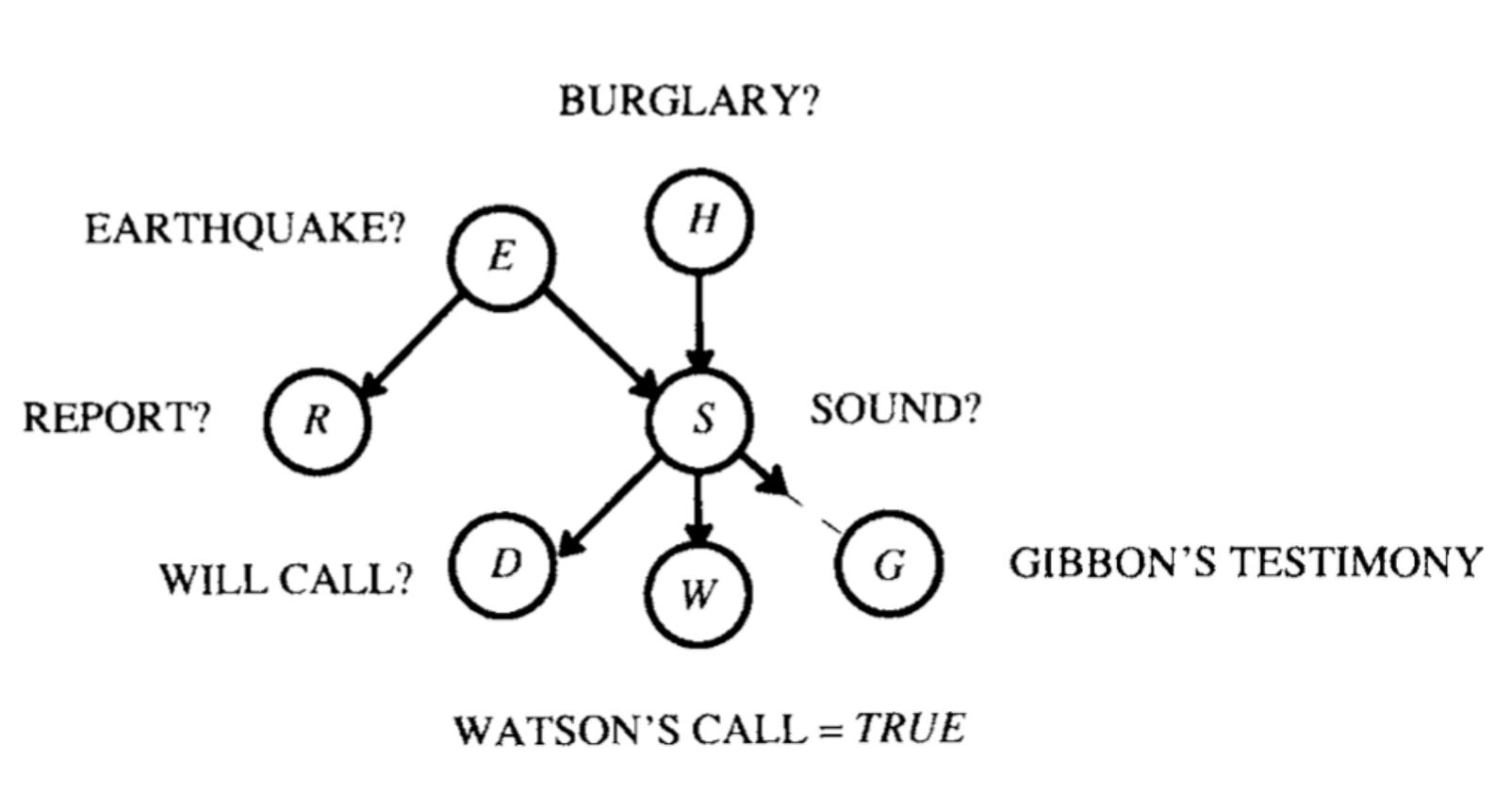
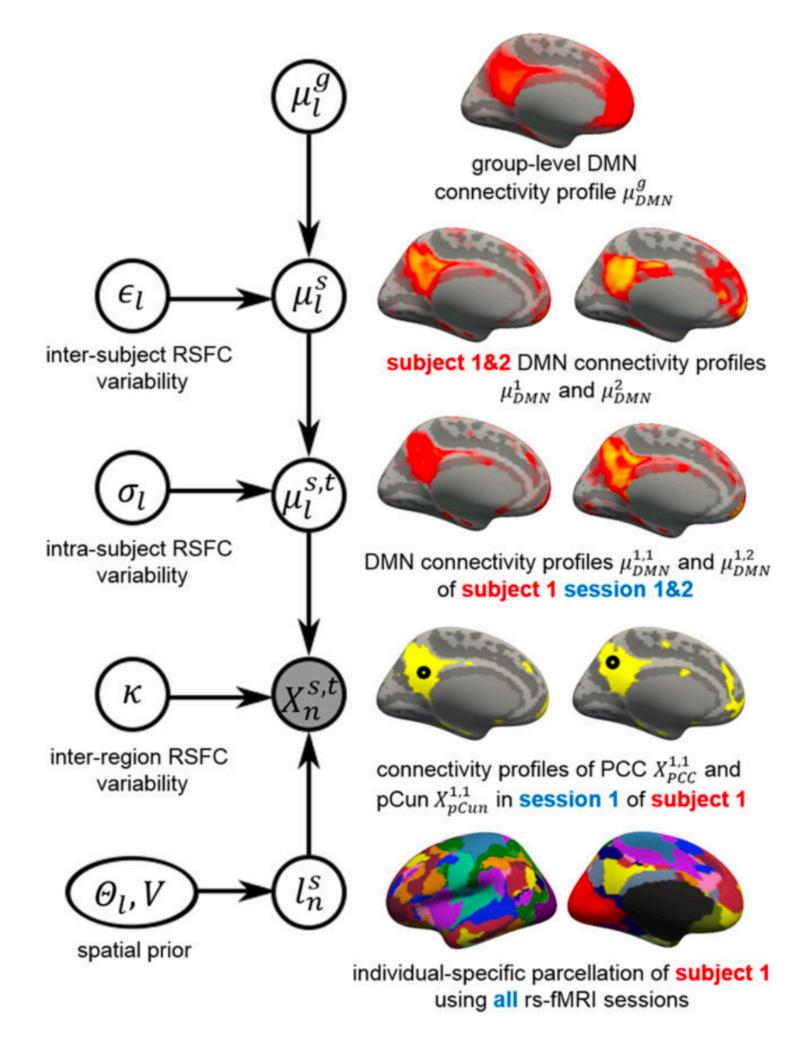
Graphical Models and Variational Inference

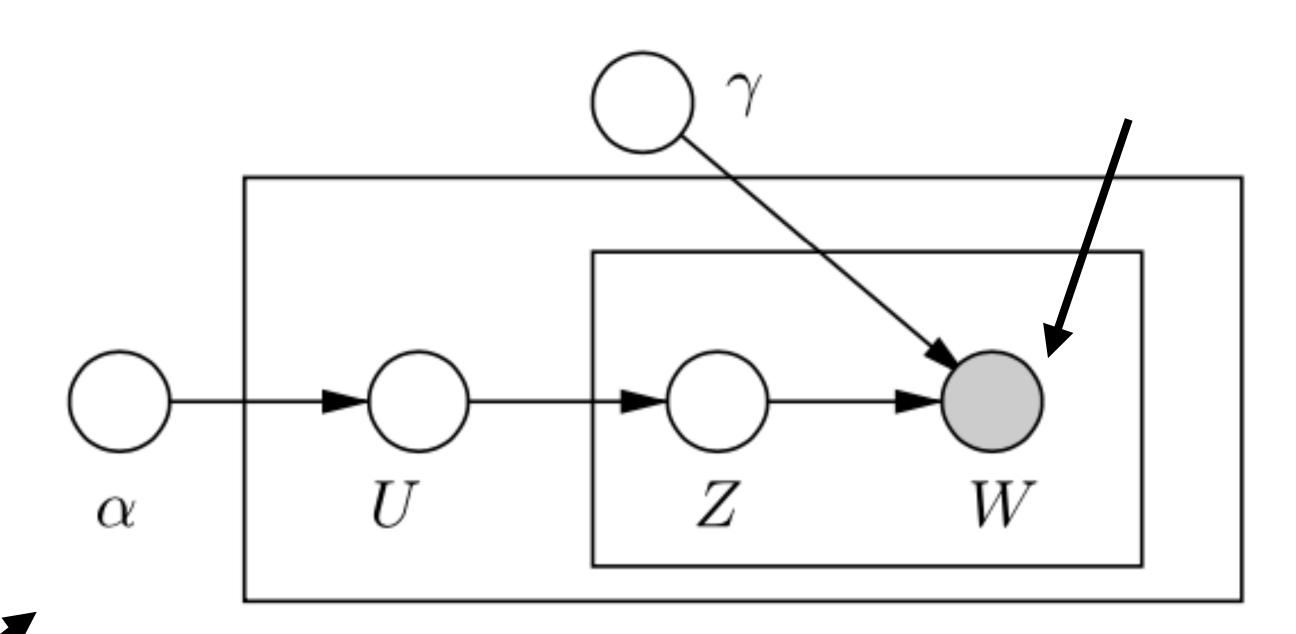
Demian Wassermann, Inria Graphical Models: Discrete Inference and Learning



[Pearl 1987]



[Kong et al 2019]



U: is a Dirichlet or "clustering variable" Z: is a "Topic"

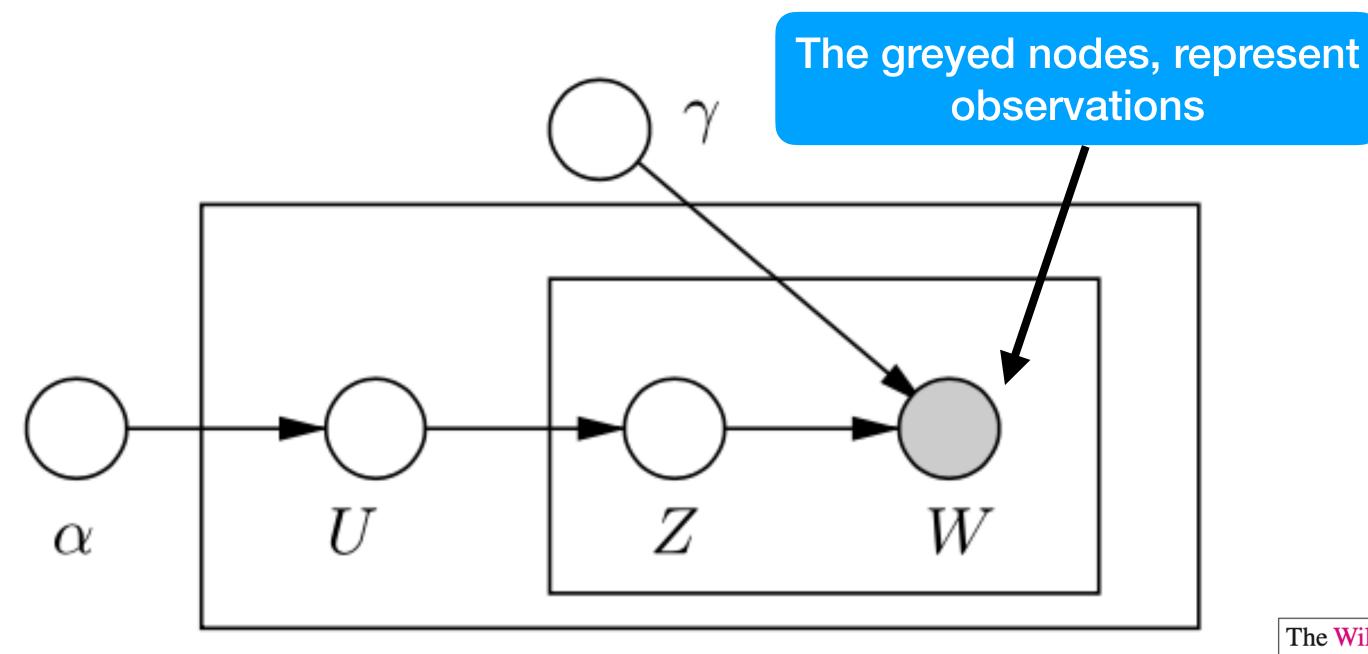
W: is an observed "Word"

[Blei et al 2003]

Each "box" or template represents a set of i.i.d. random variables with the same distribution

"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL
FILM	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
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LOVE	CONGRESS	$_{ m LIFE}$	HAITI

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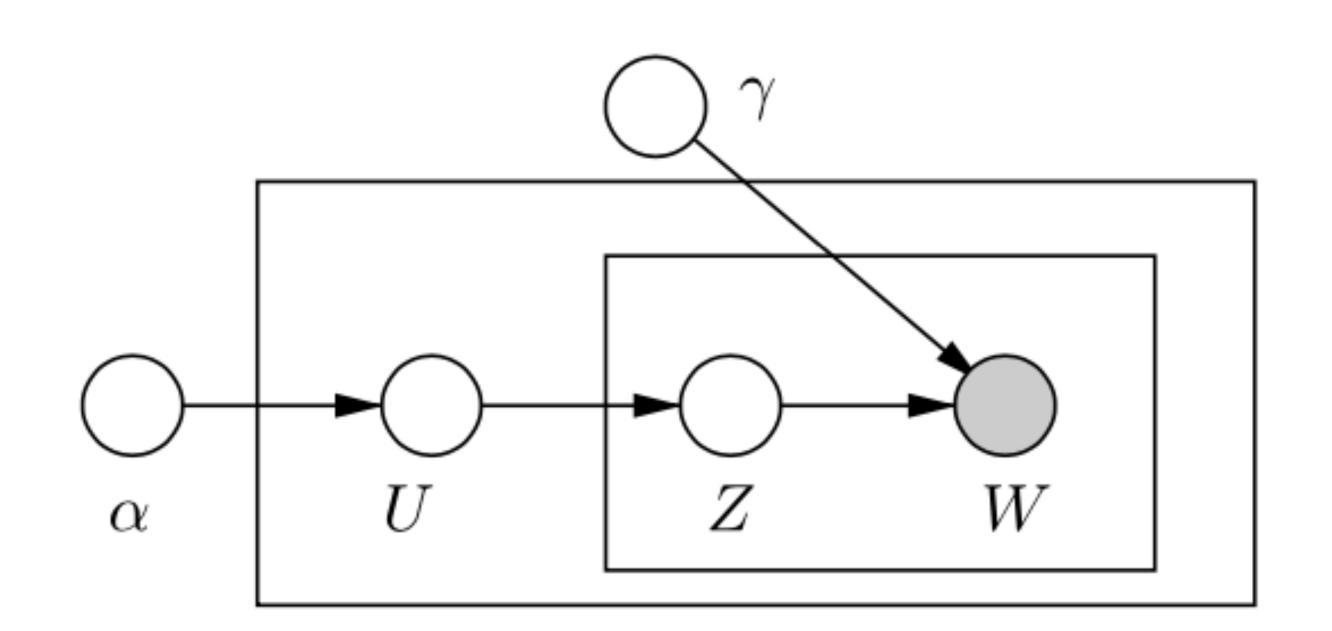
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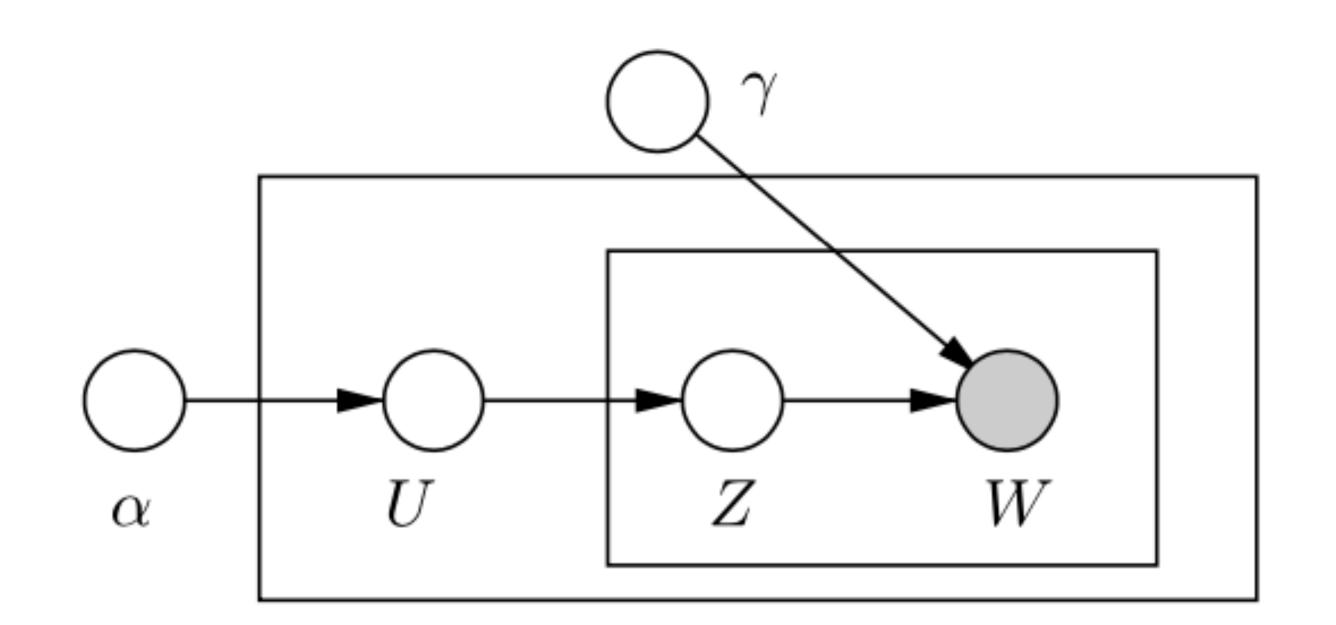
$$U_{j} \sim Dirichlet(\alpha), \alpha < 1$$

$$Z_{i,j} \sim Multinomial(U_{j})$$

$$W_{i,j} \sim Multinomial\left(\gamma_{Z_{i,j}}\right)$$

Then, we are looking for the posterior $P(U,Z|W,\alpha,\gamma) = \frac{P(U,Z,W|\alpha,\gamma)}{P(W|\alpha,\gamma)}$

$$P(U, Z, W | \alpha, \gamma) = \prod_{j} \int P(U_j | \alpha) \left(\prod_{i} \sum_{Z_{i,j}} P(Z_{i,j} | U_j) P(W_{i,j} | Z_{i,j}, \gamma) \right) dU_j$$



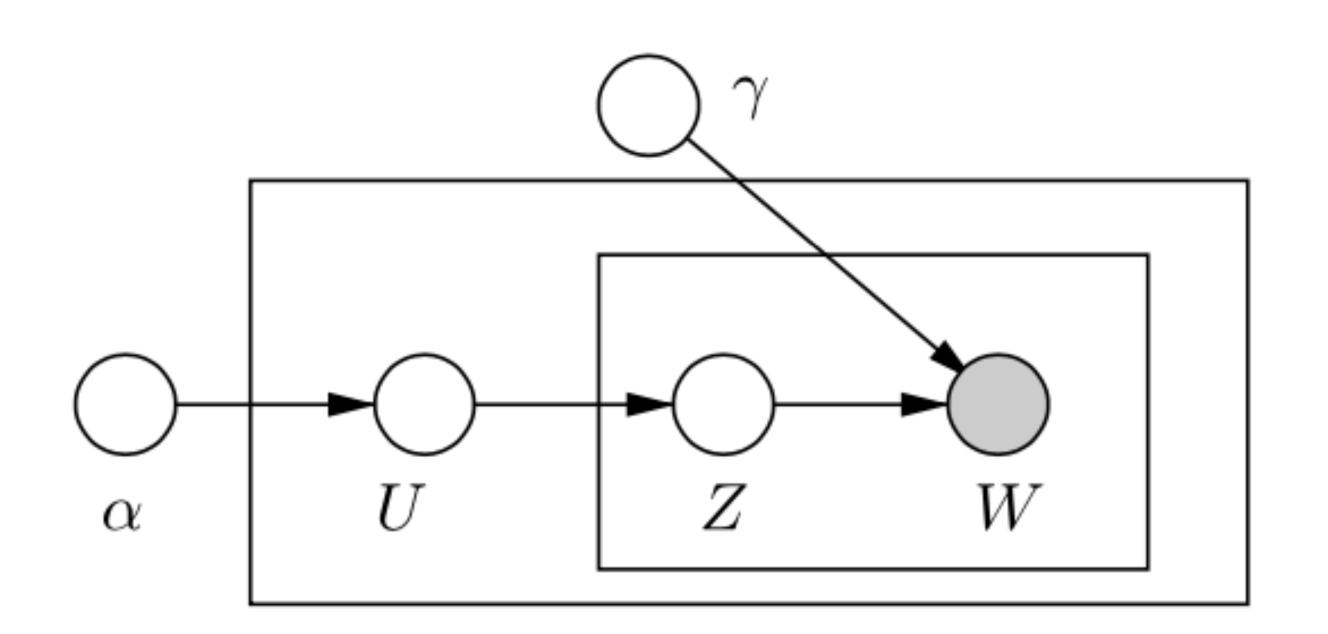
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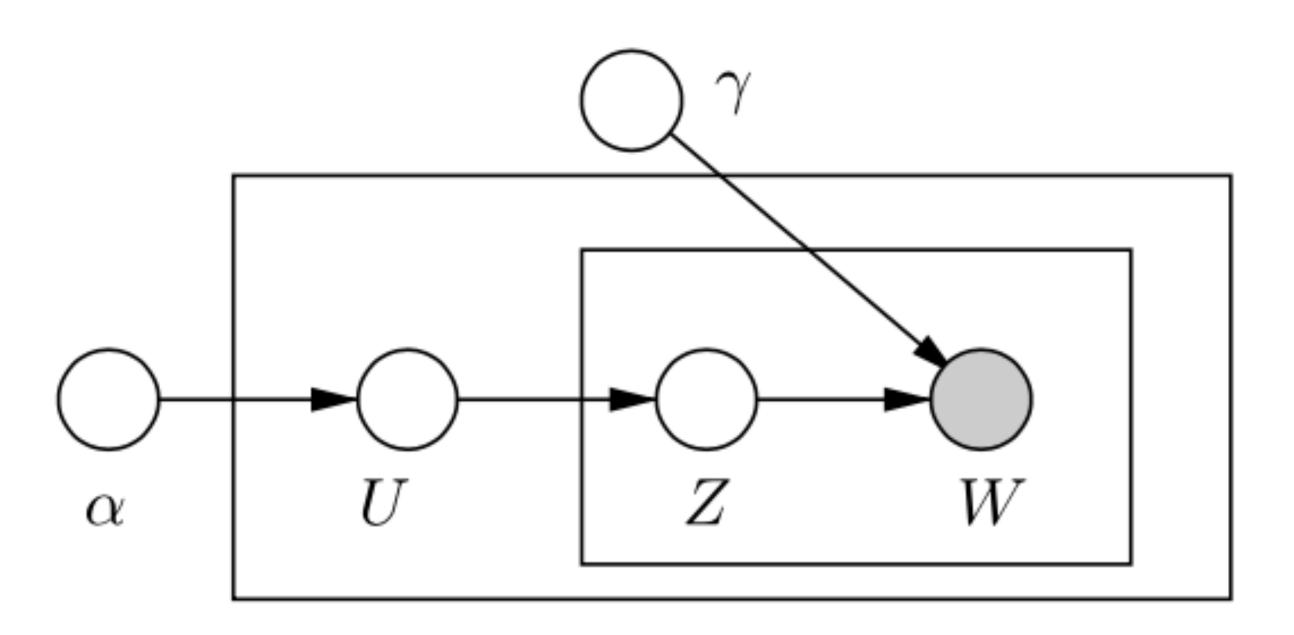
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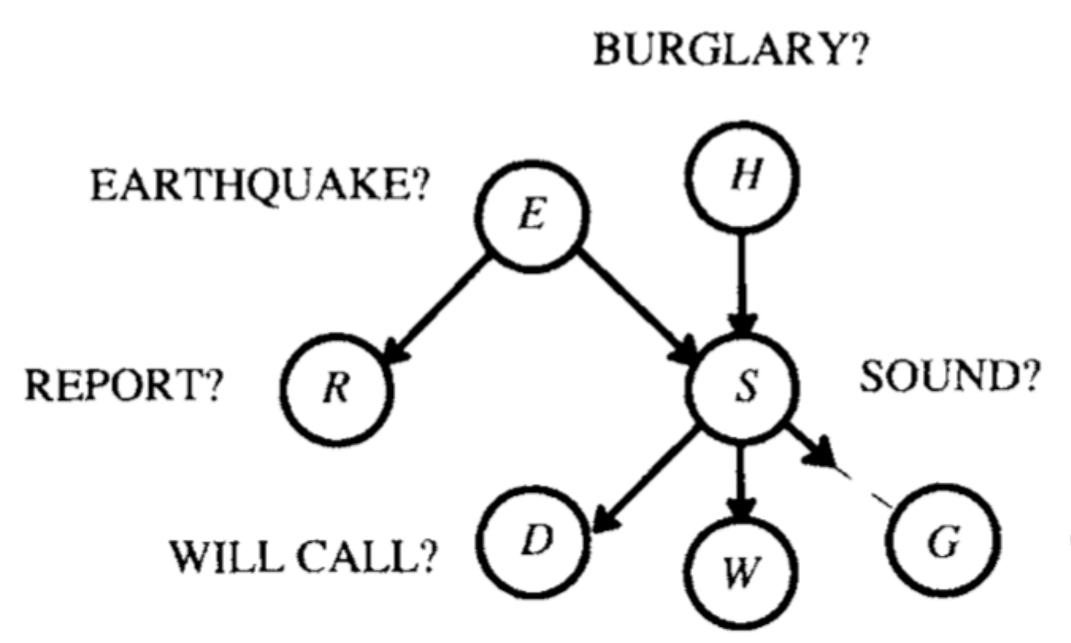
"Arts"	"Budgets"	"Children"	"Education"
"Arts" NEW FILM SHOW MUSIC MOVIE PLAY MUSICAL BEST ACTOR	"Budgets" MILLION TAX PROGRAM BUDGET BILLION FEDERAL YEAR SPENDING NEW	"Children" CHILDREN WOMEN PEOPLE CHILD YEARS FAMILIES WORK PARENTS SAYS	"Education" SCHOOL STUDENTS SCHOOLS EDUCATION TEACHERS HIGH PUBLIC TEACHER BENNETT
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$$P(W_1, ..., W_I, Z_1, ..., Z_I, U_1, ..., U_J, \alpha, \gamma) = \prod_j \prod_i P(W_i | Z_i, \gamma) P(Z_i | U_j) P(U_j | \alpha)$$

In general, for a graphical model Graphical Model with vertices V and edges E

$$GM = (V, E), P(V) = \prod_{v \in V} P(v | Pa(v)), Pa(v) = \{v' : v' \to v \in E\}$$

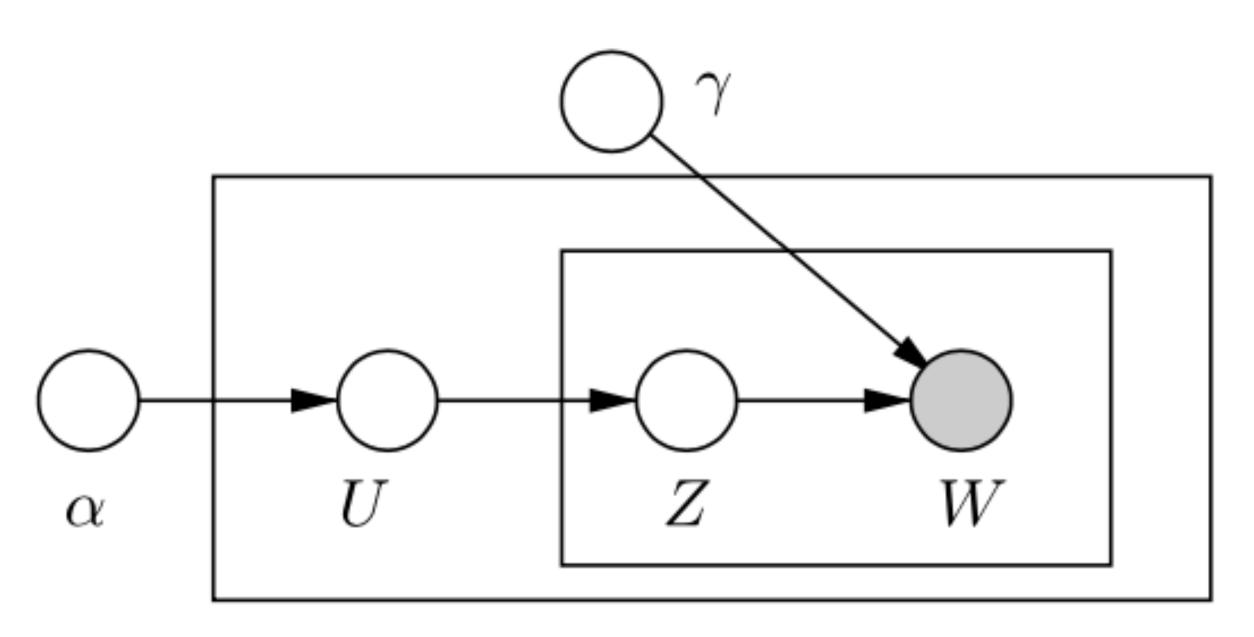


Here, the report and the sound are independent, given that we know if there was an earthquake: They are **conditionally** independent

$$P(R, S \mid E) = P(R \mid E)P(S \mid E) \text{ iif } I(R, S, E)$$

GIBBON'S TESTIMONY

WATSON'S CALL = TRUE



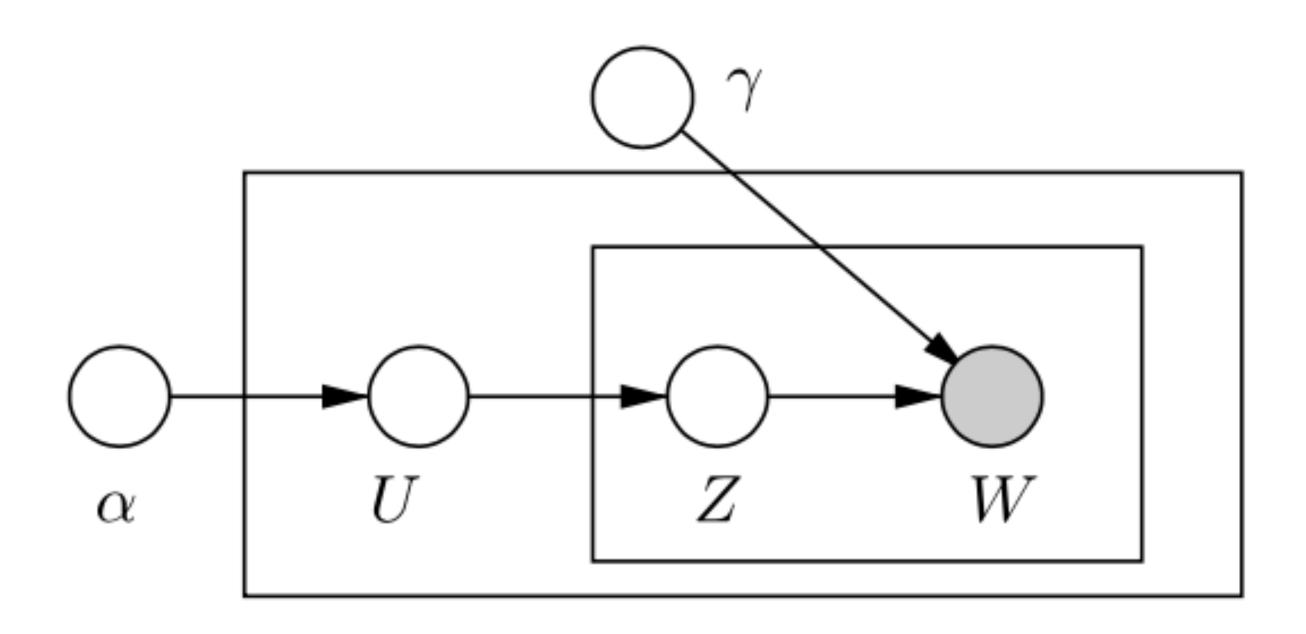
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However, our usual problem is: given observed variables O and latent variables L, to compute the posterior $P(L\,|\,O)$

$$P(L \mid O) = \frac{\prod_{v \in V} P(v \mid Pa(v))}{\prod_{o} P(o \mid Pa(o))}, GM = (V = L \cup O, E), \ \exists l \in L : o \to l \in E$$



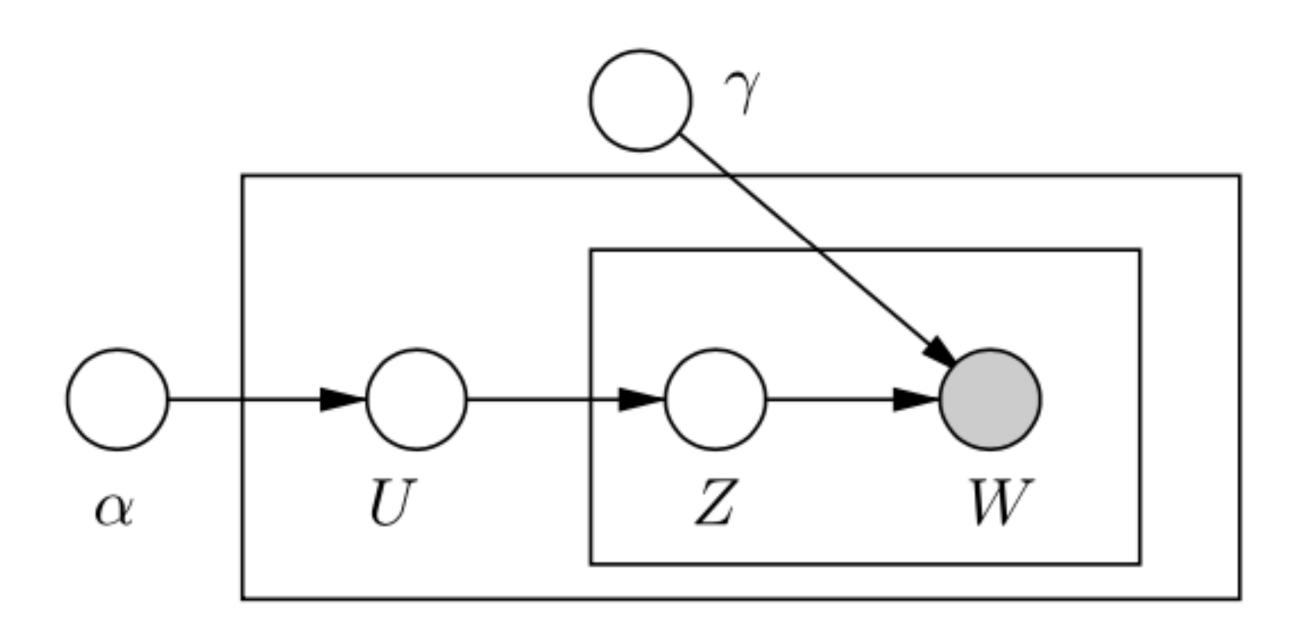
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In the case of continuous variables this is

$$P(L \mid O) = \frac{P(L, O)}{\int P(L, O)dO}$$



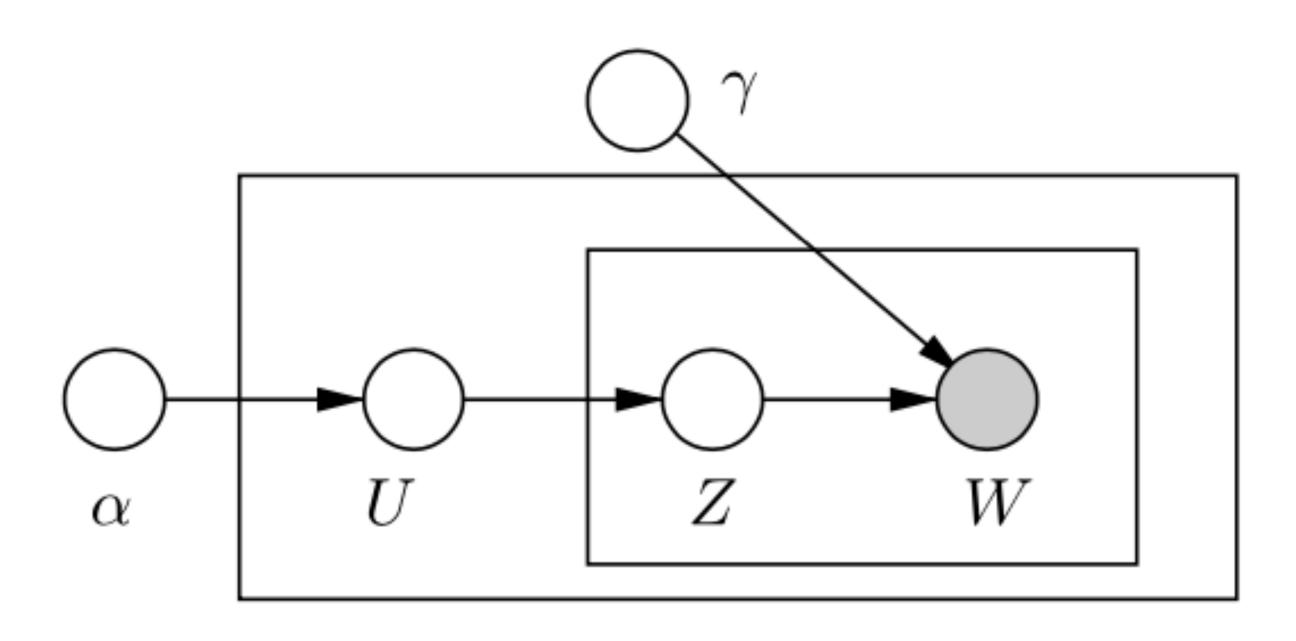
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No analytical solution, for the general case $P(L \mid O) = \frac{P(L, O)}{\int P(L, O) dO}$



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Can we approximate $P(L \mid O)$?

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- First try: MacLaurin $Q(L) = \sum P(L=l\,|\,O) + P'(L=l\,|\,O)(l-L) + \dots$ problem: how to guarantee that Q(L) is a probability law?
- Second try: cumulant approximations (changing the random L by X)

$$\phi(t) = \log \mathbb{E}_{X \sim Q(X)}[\exp(tX)] = \sum_{n} \kappa_n \frac{t^n}{n!} = \kappa_1 t + \kappa_2 \frac{t^2}{2!} + \dots = \mu t + \sigma^2 \frac{t^2}{2!} + \dots$$

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 However, a probability law has either up to two moments, or an infinite number (Cramèr 1938)

Can we approximate $P(L \mid O)$?

$$Q(L) \simeq P(L \mid O) = \frac{P(L, O)}{\int P(L, O) dO}$$

Other options: Edgesworth, approximations which come from this identity

$$\phi(t) = \log \mathbb{E}_X[\exp(itX)] = \sum_{n} \kappa_n \frac{(it)^n}{n!},$$

$$\psi(t) = \log \mathbb{E}_X[\exp(itX)] = \sum_{n=1}^{n} \gamma_n \frac{(it)^n}{n!}$$

$$\hat{\phi}(t) = \sum_{n} (\kappa_n - \gamma_n) \frac{(it)^n}{n!} + \log \psi(t)$$

however, they are not guaranteed to be probability laws for finite samples.

Can we approximate $P(L \mid O)$?

$$Q(L) \simeq P(L \mid O) = \frac{P(L, O)}{\int P(L, O) dO}$$

- So? What do we do?
 - We choose an approximate distribution $Q_{\theta}(X)$ —replacing L by X and O by Z for notation— from a given family, with parameters θ . Then $Q^* = Q_{\theta^*}: \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$

so we need to define the right similarity measurement D to compare distributions. And in standard Variational Inference (VI), Z is notation for O

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So Which D and Q Should We Choose?

$$Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$$

X the latent variables and Z the observations

Let's start with "analytical" ideas:

$$D(Q_{\theta}(X), P(X|Z)) = \int (Q_{\theta}(x) - P(x|Z))^2 dx$$

- •What does it mean for two distributions to be close in the L_2 sense?
- How easy is to obtain bounds and closed form solutions?
- $Q_{\theta}(X): X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma)$: This is called the Laplace approximation
 - •Even simpler $\Sigma=\sigma^2$ Id, which boils down to $Q_{\mu}(X)=\Pi_iQ_{\mu_i}(X_i)$

So Which D and Q Should We Choose?

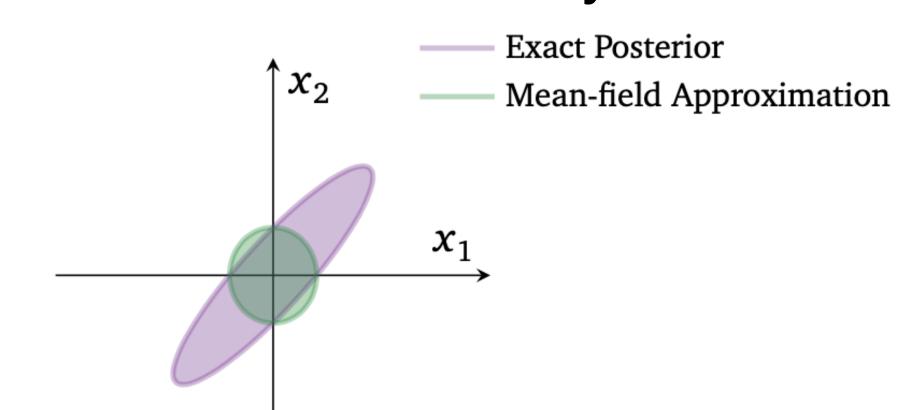
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X the latent variables and Z the observations

More Information theoretic

$$D_{KL}(Q_{\theta}(X), P(X|Z)) = \mathbb{E}_{X \sim Q_{\theta}} \left[-\log \frac{P(X|Z)}{Q_{\theta}(X)} \right] = -\int dQ_{\theta}(x) \log \frac{P(x|Z)}{Q_{\theta}(x)}$$

- The Kullback-Leibler divergence is based on information theory
- Known formulations for common cases
- -Mean field $Q_{\theta=\mu}(X)=\Pi_i Q_{\mu_i}(X_i)$



[Blei et al 2017]

A Case for Mean Field KL-based VI

Journal of Artificial Intelligence Research 4 (1996) 61—76

Submitted 11/95; published 3/96

Mean Field Theory for Sigmoid Belief Networks

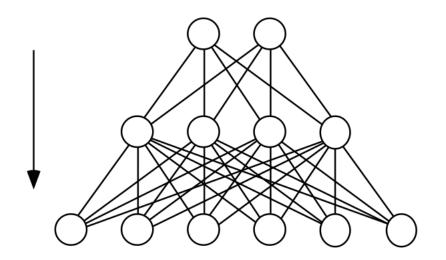
Lawrence K. Saul Tommi Jaakkola Michael I. Jordan

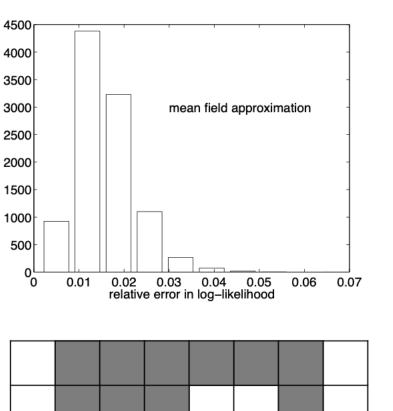
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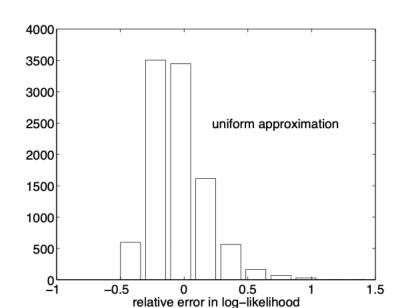
Center for Biological and Computational Learning Massachusetts Institute of Technology 79 Amherst Street, E10-243 Cambridge, MA 02139

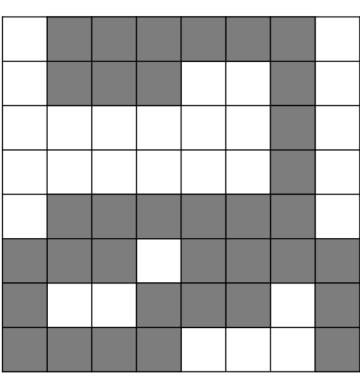
Abstract

We develop a mean field theory for sigmoid belief networks based on ideas from statistical mechanics. Our mean field theory provides a tractable approximation to the true probability distribution in these networks; it also yields a lower bound on the likelihood of evidence. We demonstrate the utility of this framework on a benchmark problem in statistical pattern recognition—the classification of handwritten digits.









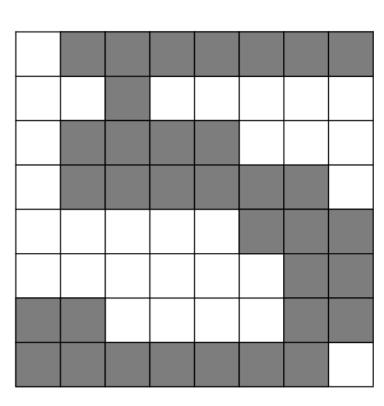


Figure 7: Binary images of handwritten digits: two and five.

	0	1	2	3	4	5	6	7	8	9
0	388	2	2	0	1	3	0	0	4	0
1	0	393	0	0	0	1	0	0	6	0
2	1	2	376	1	3	0	4	0	13	0
3	0	2	4	373	0	12	0	0	6	3
4	0	0	2	0	383	0	1	2	2	10
5	0	2	1	13	0	377	2	0	4	1
6	1	4	2	0	1	6	386	0	0	0
7	0	1	0	0	0	0	0	388	3	8
8	1	9	1	7	0	7	1	1	369	4
9	0	4	0	0	0	0	0	8	5	383

So Which D and Q Should We Choose?

$$Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$$

X the latent variables and Z the observations

A second order information-theoretic model

$$D_{KL}(Q_{\theta}(X), P(X|Z)) = \mathbb{E}_{X \sim Q_{\theta}} \left[-\log \frac{P(X|Z)}{Q_{\theta}(X)} \right] = -\int dQ_{\theta}(x) \log \frac{P(x|Z)}{Q_{\theta}(x)}$$

 $\cdot Q_{\theta}(X): X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma)$: This is called the Laplace approximation

But Laplace is Better

Journal of Machine Learning Research (2013)

Submitted 00/00; Published 00/00

Variational Inference in Nonconjugate Models

Chong Wang

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David M. Blei

Department of Computer Science Princeton University Princeton, NJ, 08540, USA

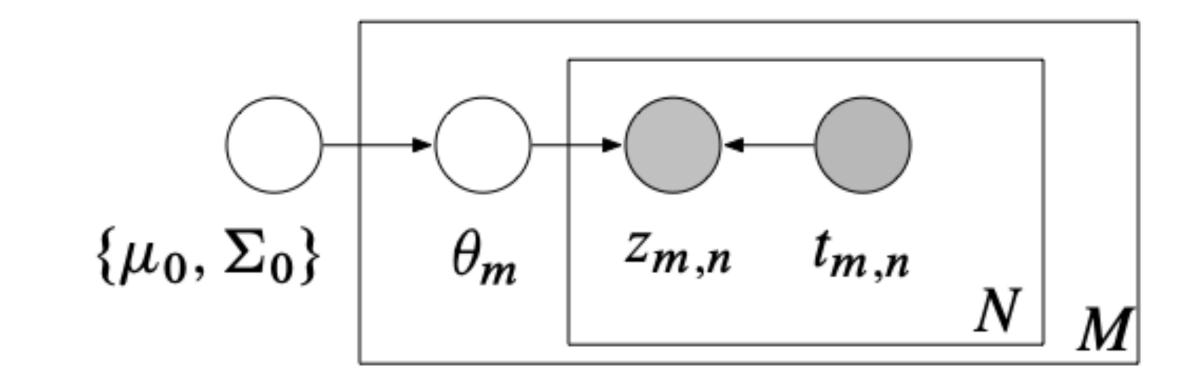








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BLEI@CS.PRINCETON.EDU

- 1. Draw coefficients $\theta \sim \mathcal{N}(\mu_0, \Sigma_0)$.
- 2. For each data point n and its covariates t_n , draw its class label from

$$z_n \mid \theta, t_n \sim \text{Bernoulli}\left(\sigma(\theta^\top t_n)^{z_{n,1}}\sigma(-\theta^\top t_n)^{z_{n,2}}\right),$$

	Yeast		Scene	
	Accuracy	Log Likelihood	Accuracy	Log Likelihood
Jaakkola and Jordan (1996)	79.7%	-0.678	87.4%	-0.670
Laplace inference	80.1%	-0.449	89.4%	-0.259

So Which D and Q Should We Choose?

$$Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$$

X the latent variables and Z the observations

A second order information-theoretic model

$$D_{KL}(Q_{\theta}(X), P(X|Z)) = \mathbb{E}_{X \sim Q_{\theta}} \left[-\log \frac{P(X|Z)}{Q_{\theta}(X)} \right] = -\int dQ_{\theta}(x) \log \frac{P(x|Z)}{Q_{\theta}(x)}$$

 $\cdot Q_{\theta}(X): X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma)$: This is called the Laplace approximation

So Which D Should We Choose? Finding Bounds

$$D_{KL}(Q_{\theta}(X), P(X|Z)) = \mathbb{E}_{X \sim Q_{\theta}} \left[-\log \frac{P(X|Z)}{Q_{\theta}(X)} \right] = -\int dQ_{\theta}(x) \log \frac{P(x|Z)}{Q_{\theta}(x)}$$

But our graphical model is more adapted to sample from P(X,Z) than from P(X|Z).

Then, can we find a way to efficiently minimise $D_{\mathit{KL}}\left(Q_{\theta}(X), \frac{P(X,Z)}{P(Z)}\right)$

when, in general, we don't know the probability of "evidence" P(Z)?

Let's see in the next slide....

So Which D Should We Choose? Finding Bounds

$$D_{KL}(Q_{\theta}(X), P(X|Z)) = \mathbb{E}_{X \sim Q_{\theta}} \left[-\log \frac{P(X|Z)}{Q_{\theta}(X)} \right] = -\int dQ_{\theta}(x) \log \frac{P(x|Z)}{Q_{\theta}(x)}$$

And we know that

$$\log P(Z) = \log \int dx P(x, Z) = \log \int \frac{dQ_{\theta}(x) P(x, Z)}{Q_{\theta}(x)} = \log \mathbb{E}_{X \sim Q_{\theta}} \left[\frac{P(X, Z)}{Q_{\theta}(X)} \right]$$

with Z being the observed data (O before) and X our latent variables (L)

then,
$$P(Z) = \log \mathbb{E}_{X \sim Q_{\theta}} \left[\frac{P(X, Z)}{Q_{\theta}(X)} \right] \geq E_{X \sim Q_{\theta}} \left[\frac{P(X, Z)}{Q_{\theta}(X)} \right] \triangleq \mathcal{L}(\theta)$$

$$\min_{\theta} D_{KL}(Q_{\theta}(X), P(X|Z)) = \log P(Z) - \max_{\theta} \mathcal{L}(\theta)$$

Hence, it is enough to maximise the Evidence Lower Bound (ELBO): $\mathscr{L}(\theta)$

So Which D and Q Should We Choose?

$$Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$$

X the latent variables and Z the observations

A simplified second order information-theoretic model

$$\theta = \arg \max_{\theta} \mathcal{L}(\theta) = \mathbb{E}_{X \sim Q_{\theta}} \left[\log \frac{P(X, Z)}{Q_{\theta}(X)} \right]$$

 ${ullet} Q_{ heta}(X): X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma)$: This is called the Laplace approximation

But Laplace is Better (they use ELBO)

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Variational Inference in Nonconjugate Models

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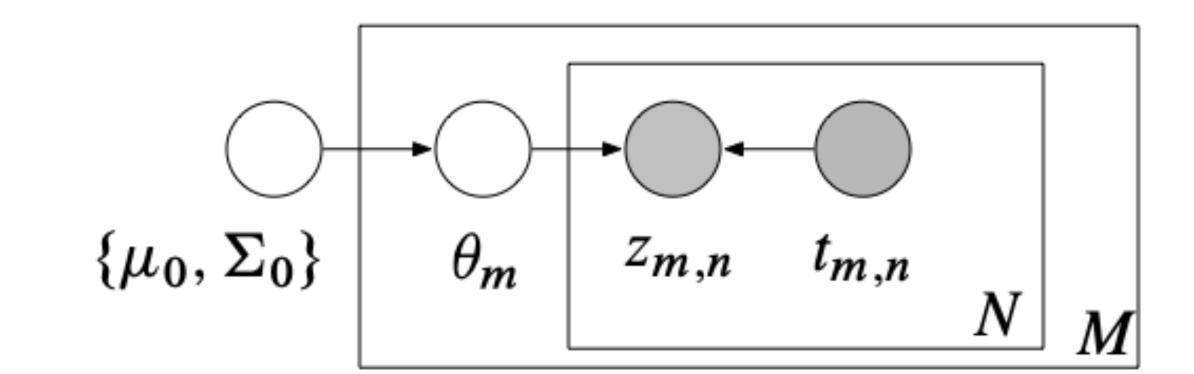








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More General Q_{θ}

$$Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$$

X the latent variables and Z the observations

•Gaussian Processes: A measure over continuous functions where any discrete sample of the domain follows a Gaussian law.

$$P(f(x)): (f(x_1), ..., f(x_N)) \sim N(\mu_{x_1,...,x_N}, \Sigma_{x_1,...,x_N})$$

•Support Transformations: $Q_{\theta}(X) \triangleq \phi_{\theta}(X)$

 $X \sim \mathcal{N}(\mu, \Sigma), \phi_{\theta}$ a parametric mass-preserving diffeomorphism

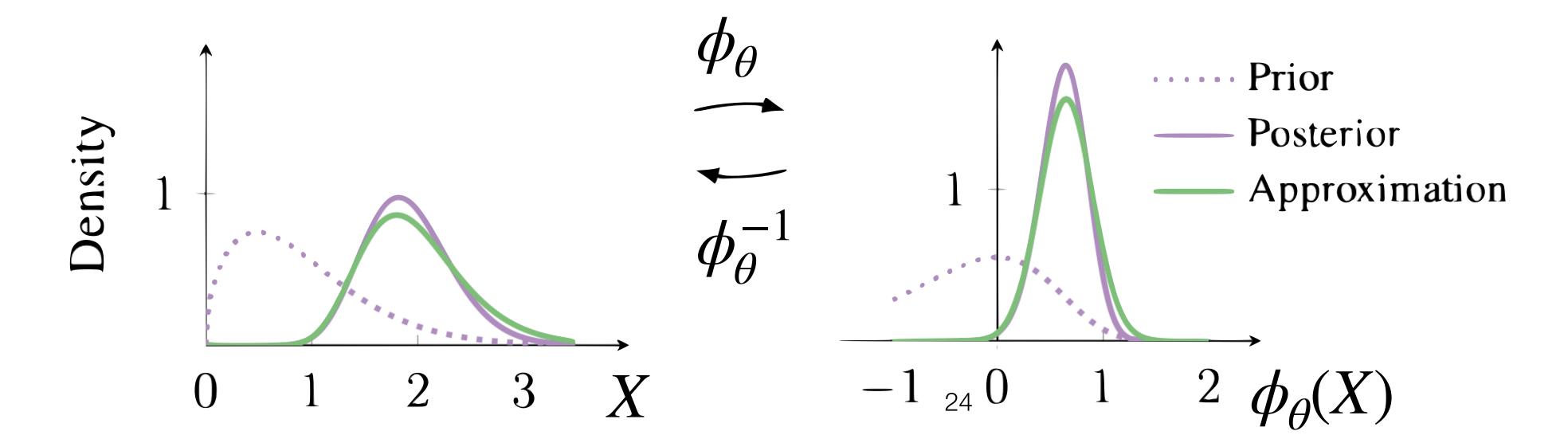
More General
$$Q_{\theta}$$

 $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$

X the latent variables and Z the observations

. Support Transformations: $Q_{\theta}(X) \triangleq N_{\mu,\Sigma}(\phi_{\theta}(X)) \left| J_{\phi_{\theta}}(X) \right|$

 $X \sim \mathcal{N}(\mu, \Sigma), \phi_{\theta}$ a parametric mass-preserving diffeomorphism



[Kucukelbir etal 17]

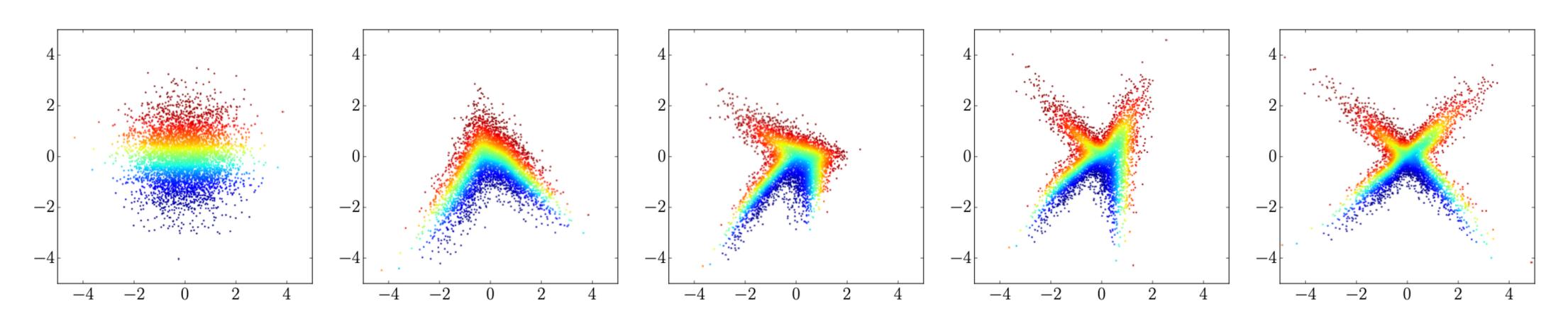
More General Q_{θ}

$$Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$$

X the latent variables and Z the observations

. Support Transformations: $Q_{\theta}(X) \triangleq N_{\mu,\Sigma}(\phi_{\theta}(X)) \left| J_{\phi_{\theta}}(X) \right|$

 $\phi_{\theta}(X) \sim \mathcal{N}(\mu, \Sigma), \phi_{\theta}$ a stochastic flow or learnable diffeomorphism



Current Problems in VI

- $(A) \rightarrow (B) \rightarrow (C)$
 - Query 1: P(B|C) = P(C|B)P(B)/P(C)
 - Query 2: $P(A|C) = \sum_{B} P(A|B)P(B|C)$
 - Amortisation, reused probability in blue

- Scalability
- Amortization [Gershman et al 2014]

Preservation of dependencies

Auto-regressive models

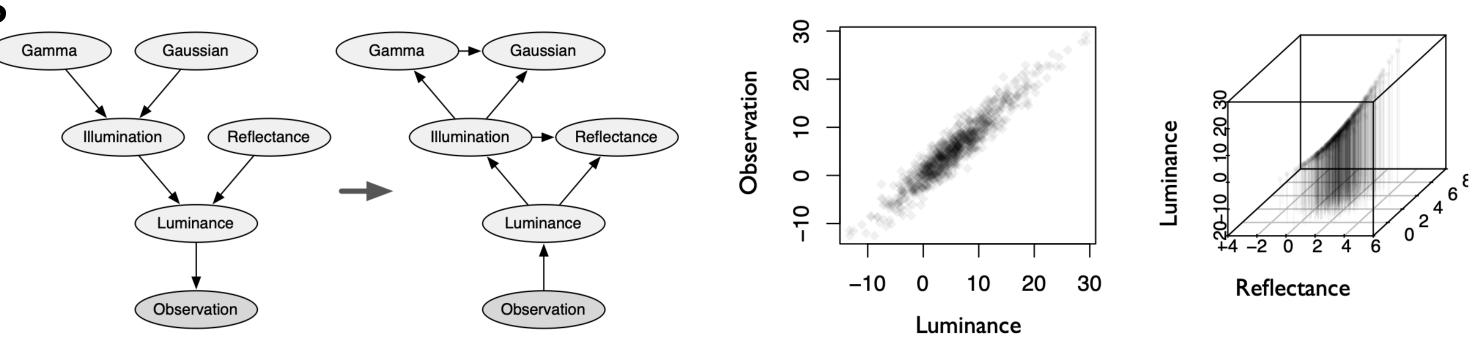
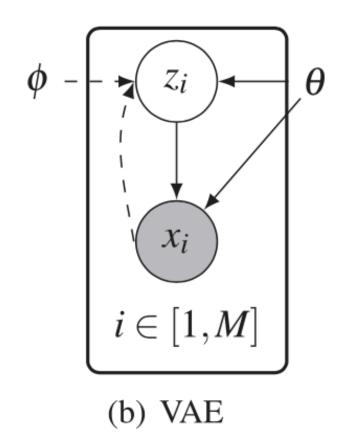


Figure 1: A Bayesian network modeling brightness constancy in visual perception, a possible inverse factorization, and two of the local joint distributions that determine the inverse conditionals.

Other Modern Bayesian Techniques

- Variational AutoEncoders
- Likelihood-free Inference

$$\theta = \arg\max_{\theta} \mathcal{L}(\theta) = \mathbb{E}_{X \sim Q_{\theta}} \left[\log \frac{P(X, Z)}{Q_{\theta}(X)} \right] \quad VAE : Z \sim N(\mu(X), \Sigma(X))$$



Likelihood Prior
$$P(Z|X) = \frac{P(X|Z)P(Z)}{P(X)}$$
Evidence